

Q<sub>1</sub> let  $0 \leq s < t$  be given, what is  $\text{Cov}[W(s), W(t)]$

$$\begin{aligned} \text{A. } W(t) &= (W(t) - W(s)) + (W(s) - W(0)) \\ &= (W(t) - W(s)) + W(s) \end{aligned} \quad \left. \vphantom{W(t)} \right\} \text{standard trick}$$

$W(t) - W(s)$  and  $W(s)$  are independent

$$W(t) - W(s) \sim N(0, t-s) \quad W(s) \sim N(0, s)$$

$$\begin{aligned} \text{Cov}[W(s), W(t)] &= \text{Cov}[W(s), W(t) - W(s) + W(s)] \\ &= \text{Cov}[W(s), W(t) - W(s)] + \text{Var}(W(s)) \\ &= E[W(s) \cdot (W(t) - W(s))] + s \\ &= E[W(s)] \cdot E[W(t) - W(s)] + s \\ &= 0 \times 0 + s \\ &= s \end{aligned}$$

Without the assumption  $s < t$ , we have the general formula  $\text{Cov}[W(s), W(t)] = \min\{s, t\}$

Q<sub>2</sub> let  $0 \leq s < t$  be given. what is  $E[W(t) | W(s) = x]$

$$\text{A}_2: E[W(t) | W(s) = x]$$

$$= E[W(t) - W(s) + W(s) | W(s) = x]$$

$$= E[W(t) - W(s) | W(s) = x] + E[W(s) | W(s) = x]$$

$$= E[W(t) - W(s)] + x$$

$$= 0 + x = x$$

martingale property of  
Brownian motion

Q3 Let  $0 \leq s < t$  be given, what is  $E[W(s) | W(t) = x]$

A3:  $W(s) = \alpha W(t) + (W(s) - \alpha W(t))$

choose  $\alpha$  so that  $\alpha W(t)$  and  $W(s) - \alpha W(t)$  are independent

$$\begin{aligned} \text{Then } E[W(s) | W(t) = x] &= E[\alpha W(t) + W(s) - \alpha W(t) | W(t) = x] \\ &= \alpha x \end{aligned}$$

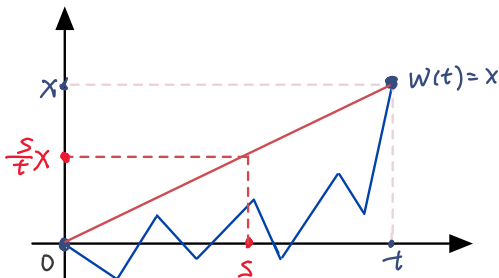
to find  $\alpha$ , we have

$$\begin{aligned} &\text{Cov}[W(t), W(s) - \alpha W(t)] \\ &= E[W(t)(W(s) - \alpha W(t))] = E[W(t)W(s)] - \alpha E[W^2(t)] \\ &= t \cdot s - \alpha \cdot t^2 = 0 \end{aligned}$$

$$\text{thus } \alpha = \frac{s}{t}$$

$$\therefore E[W(s) | W(t) = x] = \frac{s}{t}x$$

to remember the result, use follow plot

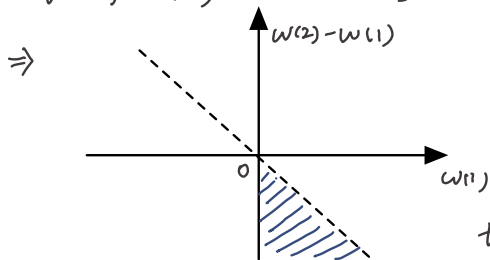


Q4: what is  $P\{W(1) > 0 \text{ and } W(2) < 0\}$ ?

A4:  $W(1) > 0$  and  $W(2) < 0$

$$\Leftrightarrow W(1) > 0 \text{ and } W(2) - W(1) + W(1) < 0$$

$W(2) - W(1)$  and  $W(1)$  are independent



joint distribution of  $W(2) - W(1)$  and  $W(1)$  is bivariate normal

$$\text{thus } P\{W(1) > 0 \text{ and } W(2) < 0\} = \frac{1}{8}$$

Q5 Is  $X(t) = \cosh(\lambda W(t)) e^{-\frac{1}{2}\lambda^2 t}$  a martingale?

$$(\cosh(x) = \frac{1}{2}(e^x + e^{-x}))$$

$$\begin{aligned} \text{A5: } X(t) &= \frac{1}{2} (e^{\lambda W(t)} + e^{-\lambda W(t)}) e^{-\frac{1}{2}\lambda^2 t} \\ &= \frac{1}{2} \left( \underbrace{e^{\lambda W(t) - \frac{1}{2}\lambda^2 t}}_{\text{martingale}} + \underbrace{e^{-\lambda W(t) - \frac{1}{2}\lambda^2 t}}_{\text{martingale}} \right) \\ &= \text{Martingale} \end{aligned}$$

Q6: Is  $2^{W(t)}$  a martingale?

A6:  $M(t) = e^{(\log 2) W(t) - \frac{1}{2}(\log 2)^2 t}$

is a martingale

but  $e^{(\log 2) W(t)} = 2^{W(t)}$  is not

Q7: Let  $S(t) = S(0) \exp[\sigma W(t) + (\alpha - \frac{1}{2}\sigma^2)t]$

what is  $\lim_{t \rightarrow \infty} W(t)$ ,  $\lim_{t \rightarrow \infty} \frac{W(t)}{t}$ ,  $\lim_{t \rightarrow \infty} S(t)$ ,  $\lim_{t \rightarrow \infty} E[S(t)]$

A7:  $\triangleright$  for  $W(t)$ , as  $t \rightarrow \infty$ , the paths of  $W(t)$  oscillate ever more widely, thus  $\lim_{t \rightarrow \infty} W(t)$  does not exist

$\triangleright$  for  $\lim_{t \rightarrow \infty} \frac{W(t)}{t}$  as  $t \rightarrow \infty$

$$E\left[\frac{1}{t} W(t)\right] = 0 \quad \text{Var}\left[\frac{1}{t} W(t)\right] = \frac{1}{t^2} \text{Var}(W(t)) = \frac{1}{t}$$

the randomness in  $\frac{1}{t} W(t)$  disappears because its variance converges to zero

thus  $\lim_{t \rightarrow \infty} \frac{W(t)}{t} = 0$

$\triangleright$  for  $S(t)$

if  $\alpha - \frac{1}{2}\sigma^2 > 0$

positive limit

$$S(t) = S(0) \exp\left[t\left(\frac{\sigma}{t} W(t) - (\alpha - \frac{1}{2}\sigma^2)\right)\right] \rightarrow \infty \text{ as } t \rightarrow \infty$$

if  $\alpha - \frac{1}{2}\sigma^2 < 0$

$$S(t) = S(0) \exp \left[ t \underbrace{\left[ -\frac{\sigma}{t} W(t) - \left( \alpha - \frac{1}{2}\sigma^2 \right) \right]}_{\text{negative limit}} \right] \rightarrow 0 \text{ as } t \rightarrow \infty$$

if  $\alpha - \frac{1}{2}\sigma^2 = 0$

$S(t) = S(0) e^{\sigma W(t)}$  does not have a limit because  $W(t)$  does not have a limit when  $t \rightarrow \infty$

▷ for  $E[S(t)]$

$$= S(0) e^{\alpha t}$$

if  $\alpha > 0$   $E[S(t)] \rightarrow \infty$  as  $t \rightarrow \infty$

if  $\alpha < 0$   $E[S(t)] \rightarrow 0$  as  $t \rightarrow \infty$

if  $\alpha = 0$   $E[S(t)] = S(0) \forall t \geq 0$

Q8: Show that when  $0 \leq s < t$

$$E \left[ \left( W(s) - \frac{s}{t} W(t) \right)^2 \mid W(t) = x \right] = \frac{s(t-s)}{t}$$

A8:  $W(s) - \frac{s}{t} W(t)$  is independent of  $W(t)$

so does  $\left( W(s) - \frac{s}{t} W(t) \right)^2$

$$\therefore E \left[ \left( W(s) - \frac{s}{t} W(t) \right)^2 \mid W(t) = x \right] = E \left[ \left( W(s) - \frac{s}{t} W(t) \right)^2 \right]$$

$$= E \left[ W^2(s) - \frac{2s}{t} W(s) W(t) + \frac{s^2}{t^2} W^2(t) \right]$$

$$= s + \frac{s^2}{t} - \frac{2s}{t} E[W(s)W(t)]$$

$$\begin{aligned} E[W(s)W(t)] &= E[(W(t)-W(s)+W(s))W(s)] \\ &= E[W(s)(W(t)-W(s)) + W^2(s)] \\ &= s \end{aligned}$$

$$\begin{aligned} \therefore E\left[\left(W(s) - \frac{s}{t}x\right)^2 \mid W(t)=x\right] &= s + \frac{s^2}{t} - \frac{2s}{t} \\ &= s - \frac{s^2}{t} \\ &= \frac{s(t-s)}{t} \end{aligned}$$