

§ Chapter 3. Conditioning, Martingales, Stopping times

Recall: (Ω, \mathcal{F}, P) , X is \mathcal{F} -measurable

$$\Leftrightarrow \{w \in \Omega : X(w) \leq c\} \in \mathcal{F} \quad \forall c \in \mathbb{R}$$

Recall Conditional probability

Take $B \in \mathcal{F}$, $P(B) > 0$, Then the probability of an event A given B is $P(A|B) = \frac{P(A \cap B)}{P(B)}$

Example: $A = \{\text{heads on first throw}\} = \{\text{HH, HT}\}$

$B = \{\text{heads at least once}\} = \{\text{HH, HT, TH}\}$

$$P(A) = P(\{\text{HH, HT}\}) = P(\{\text{HH}\}) + P(\{\text{HT}\}) = \frac{1}{4} + \frac{1}{4} = \frac{1}{2}$$

$$P(A \cap B) = P(\{\text{HH, HT}\}) = \frac{1}{2}$$

$$P(B) = P(\{\text{HH, HT, TH}\}) = \frac{3}{4}$$

$$\Rightarrow P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{2}{3}$$

Two questions:
 (motivation) \Leftrightarrow How to condition on a σ -algebra \mathcal{G} rather than a specific event B (i.e. $P(A|\mathcal{G})$)

\Leftrightarrow How to condition on events for which

$$P(B) = 0$$

Extensions:

\Leftrightarrow Assume $\mathcal{G} = \sigma(B_1, B_2, \dots, B_N) \quad n \in \mathbb{N} \quad B_i \cap B_j = \emptyset \quad \forall i \neq j$

$$\bigcup_{i=1}^N B_i = \Omega$$

N can be extended
 \uparrow to ∞

$$1_{B_i}(w) = \begin{cases} 1, & w \in B_i \\ 0, & w \notin B_i \end{cases}$$

Define $P(A|\mathcal{G})(w) := \sum_{i=1}^N P(A|B_i) 1_{B_i}(w)$

$$\text{Ex. } w \in B_6 \quad 1_{B_i}(w) = \begin{cases} 1 & i=6 \\ 0 & \text{otherwise} \end{cases} \Rightarrow P(A|G)(w) = P(A|B_6)$$

More generally, take a r.v. X , then we define the conditional expectation

$$E[X|G](w) = \sum_{i=1}^{\infty} E[X_i|B_i] \cdot 1_{B_i}(w)$$

Example: coin-toss . $B_1 = \{HT, HH\}$, $B_2 = \{TH, TT\}$

$$G = \sigma(B_1, B_2) = \{\emptyset, \phi, B_1, B_2\}$$

$X \triangleq$ number of heads

$$\begin{aligned} E[X|G] &= E[X|B_1] \cdot 1_{B_1} + E[X|B_2] \cdot 1_{B_2} \\ &= \frac{1+2}{2} 1_{B_1} + \frac{0+1}{2} 1_{B_2} \\ &= \frac{3}{2} 1_{B_1} + \frac{1}{2} 1_{B_2} \end{aligned}$$

<2> Def: (Sub- σ -algebra) if G is a σ -algebra, and

$$A \in G \Rightarrow A \in F \quad \forall A \in G$$

Def 3.0.1
(Conditional Expectation) Let X be a (F -measurable) r.v. $G \subseteq F$ is a Sub- σ -algebra. We define $E[X|G]$, the conditional

only information expectation of X given G , to be a r.v. satisfying contained in

g is used to compute conditional expectation ① $E[X|G]$ is G -measurable (not F -measurable in general)

② $\forall G \in G, E[X|G] = E[\underbrace{E[X|G]}_{\substack{\text{r.v.} \\ \text{F-measurable.}}} \cdot 1_G]$ (averaging identity)

not G -measurable. \downarrow in general, this will not hold for events $F \in F/G$

Remark:

$\Delta G \subseteq F: \Rightarrow G$ contains less information than F .

- ▷ $E[X|G]$ is the "best" approximation of X on the subsets of G .
 - ▷ Conditional expectation is unique. For any G -measurable r.v. y satisfies $E[X1_G] = E[Y1_G] \forall G \in \mathcal{G}$, we have

$$y = E[X|G]$$

\Rightarrow we often "guess" $E[X|G]$, and then check (1) & (2) of the definition
 - ▷ Plug in $G = \Omega$ into ②
- $$E[X] = E[X1_\Omega] = E[E[X|G]1_\Omega] = E[E[X|G]]$$
- \Rightarrow law of total probability
- ▷ assume $G = \sigma(Y)$ for another r.v. y on (Ω, \mathcal{F}, P)
- $$E[X|Y] := E[X|G] = g(y)$$
- Lemma: A $\sigma(Y)$ -measurable r.v. Z can be written as a function of y , i.e. $g: \mathbb{R} \rightarrow \mathbb{R}$ $Z = g(Y)$

Property 3.0.2

- ① Linearity: If X, Y are r.v and $\lambda \in \mathbb{R}$, then

$$E[X + \lambda Y|G] = E[X|G] + \lambda E[Y|G]$$
- ② Positivity: If $X \leq Y$, then $E[X|G] \leq E[Y|G]$
- ③ "Take-out-what's-known"
 If X is G -measurable and Y is any r.v. then

$$E[XY|G] = X E[Y|G]$$

④ Tower Property: If $H \subseteq G \subseteq F$ are σ -algebras, then

$$\begin{aligned} E[X|H] &= E[E[X|G]|H] \quad (\text{the smaller } \sigma\text{-algebra always wins}) \\ &= E[E[X|H]|G] \end{aligned}$$

Remark:

$G \subseteq F$, X is F -measurable, Y is G -measurable, then

$$E[X|Y] = E[E[X|Y|G]] = E[Y|E[X|G]]$$

Property 3.0.1 If X is G -measurable, then $E[X|G] = X|E[1|G]$

If X is independent of G , that is

$$P(\{X \leq t\} \cap G) = P(X \leq t) \cdot P(G) \quad \forall t \in \mathbb{R}, G \in \mathcal{G}$$

then $E[X|G] = E[X]$

In particular, If $G = \sigma(Y)$ and X and Y are independent, then

$$E[X|Y] = E[X|\sigma(Y)] = E[X]$$

Lemma 3.0.1 (Independence Lemma). Suppose X, Y are two random variables such that X is independent of G and Y is G -measurable. Then if $g = g(x, y)$ is any function of two variables we have

$$E[g(x, Y)|G] = h(Y)$$

where $h = h(y)$ is the function defined by

$$h(y) := E[g(x, y)]$$

3.1. Martingales

Example: Assume today is time t , and you want to invest until $T \geq t$. Your portfolio has price $(X_s)_{t \leq s \leq T}$

If X is a martingale, then "on average", you will not be able to lose or win money on your investment in the future

$$\rightarrow E[\underbrace{X_T}_{\substack{\downarrow \\ \text{future value of the portfolio}}}|F_t] = X_t \quad (T \geq t)$$

"Martingales are fair games"

3.2. Adapted processes & filtration

Def 3.2.1: A filtration is a family of σ -algebra $(F_t)_{t \geq 0}$ such that whenever $s \leq t$ we have $F_s \subseteq F_t$

Def 3.2.3. Given a stochastic process X , the filtration generated by X is the family of σ -algebra $(F_t^X)_{t \geq 0}$ so that $\forall t \quad (X_{s_1}, X_{s_2}, \dots) \quad s_1, s_2, \dots \leq t$ is F_t^X -measurable i.e. F_t^X is generated by all events that can be observed using $\{X_s : s \leq t\}$

Def 3.2.4. (Adapted process). A stochastic process X is said to be adapted to $(F_t)_{t \geq 0}$ if for all $t \geq 0$, X_t is F_t -measurable

$\{X_t < c_i\} \in F_t \quad \forall c_i \in \mathbb{R}$

(definition of measurable)

Remark:

X is always adapted to $(\mathcal{F}_t^x)_{t \geq 0}$

(recall that a r.v. y is $\mathcal{G}(y)$ -measurable)

Def 3.3.1 (Martingale) A stochastic process H is a martingale ^(MG)

w.r.t. $(\mathcal{F}_t)_{t \geq 0}$ if

① H is adapted to $(\mathcal{F}_t)_{t \geq 0}$

② $\forall s \leq t$, we have $E[M_t | \mathcal{F}_s] = M_s$ (fair game)

Def 3.3.2 Sub-Martingale: \rightarrow "earn money on average"

cond ① & $E[M_t | \mathcal{F}_s] \geq M_s \quad \forall t \geq s$

Super-Martingale:

\rightarrow "losing money on average"

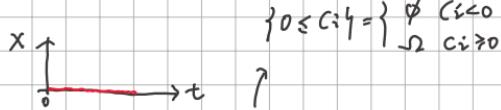
cond ① & $E[M_t | \mathcal{F}_s] \leq M_s \quad \forall t \geq s$

Example:

• Martingale

$$M_t = D \quad \forall t \geq 0$$

$$E[M_t | \mathcal{F}_s] = M_s \quad \forall t \geq s$$



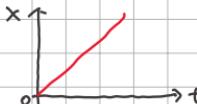
$$\{0 \leq c_i \cdot 1 = \} \not\subset \{c_i \geq 0\}$$

$$\{M_t \leq c_i \cdot 1 \in \mathcal{F}_t \quad \forall c_i \in \mathbb{R}\}$$

• Sub-martingale

$$M_t = t \quad \forall t \geq 0$$

$$E[M_t | \mathcal{F}_s] = t \geq s = M_s \quad \forall t \geq s$$



$\Rightarrow H$ is adapted
to $(\mathcal{F}_t)_{t \geq 0}$

• Super-martingale

$$M_t = -t$$

Example Let $T > 0$ and $(F_t)_{0 \leq t \leq T}$, X is F_T -measurable

Then $X_t := E[X|F_t]$ is a martingale w.r.t. $(F_t)_{0 \leq t \leq T}$

proof: $X_t = E[X|F_t] \Rightarrow X_t$ is adapted to $(F_t)_{t \geq 0}$

$$\begin{aligned} \forall s \geq t \quad E[X_s | F_t] &= E[E[X | F_s] | F_t] \\ &= E[X | F_t] = X_t \end{aligned}$$

therefore X_t is a martingale w.r.t. $(F_t)_{0 \leq t \leq T}$

3.4. Martingale property of random walks (discrete times)

In Lecture 1, we defined $S_n = \sum_{n=1}^k \xi_n$

Take $F_k = \mathcal{F}_k^S$

<1> S is $(\mathcal{F}_k^S)_{k \geq 0}$ - adapted by definition

<2> Take $k, l \in \mathbb{N}$ and $k \leq l$

$$E[S_l | F_k] = E[S_l - S_k + S_k | F_k]$$

$\rightarrow S_k$ is adapted to F_k

$$(\text{linearity}) = E[S_l - S_k | F_k] + E[S_k | F_k]$$

$$= E\left[\sum_{n=k+1}^l \xi_n | F_k\right] + S_k$$

$$(\text{linearity}) = \sum_{n=k+1}^l E[\xi_n | F_k] + S_k$$

$$= \sum_{n=k+1}^l E[\xi_n] + S_k \quad (\xi_{k+1}, \dots, \xi_l) \text{ is indep. of } (\xi_1, \dots, \xi_k)$$

$$= S_k$$

3.5. Martingale property of Brownian motion

Theorem 3.5.1 Let W be a Brownian motion, $F_t = \mathcal{F}_t^W$, then

W is a MG w.r.t. $(F_t)_{t \geq 0}$

Proof: Recall independence of increments

$$\forall 0 \leq t_1 \leq \dots \leq t_n$$

$W_{t_1}, W_{t_2} - W_{t_1}, \dots, W_{t_n} - W_{t_{n-1}}$ are independent r.v.s
 $(t \geq s)$

In particular, $W_t - W_s$ is independent of \mathcal{F}_s . $\forall r \in s$

$\Rightarrow W_t - W_s$ is independent of \mathcal{F}_s^W

$$\text{Thus } E[W_t | \mathcal{F}_s] = E[W_t - W_s + W_s | \mathcal{F}_s]$$

$$= E[W_t - W_s | \mathcal{F}_s] + E[W_s | \mathcal{F}_s]$$

$$= \underbrace{E[W_t - W_s]}_{\sim N(0, (t-s))} + W_s$$

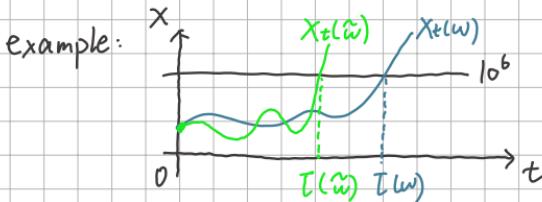
$$= 0 + W_s = W_s$$

3.6. Stopping times

Imagine you want to decide when to stop investing depending on your current & historical walk. Your stopping rule is given by a random time, which depends on the information available to you at a specific time.

Def 3.6.1 (stopping time). A stopping time is a function

$$T: \Omega \rightarrow [0, \infty), \text{ such that } \{T \leq t\} \in \mathcal{F}_t \quad \forall t \geq 0$$



Prop. 3.6.1 Let X be a continuous adapted process, let $\alpha \in \mathbb{R}$

$$T := \arg \min_t \{t \geq 0 : X_t = \alpha\}. \text{ Thus } T \text{ is a stopping time.}$$

It is called the (first) hitting time of level α .

Theorem 3.6.1. (Doob's optional sampling theorem)

Let M be a martingale, T be a ^{bounded} stopping time,

then the stopped process $M_T^T := M_{T \wedge T}$ is also a martingale.

Consequently, $E[M_T^T] = E[M_{T \wedge T}] = E[M_0] = E[M_T]$ $\forall t \geq 0$

(the proof is omitted)

Remark:

if instead of assuming T is bounded, we assume M^T is bounded, the Doob's optional sampling theorem is still valid.

One failed case for this theorem, for example, Let W be a standard Brownian motion (that is $W_0 = 0$ and $\text{Var}(W_t) = t$) let T be the first hitting time of W to 1. Then obviously $E[W_T] = 1 \neq 0 = E[W_0]$. This is because $E[T] = \infty$ (or $\exists \omega \in \Omega$ $T(\omega) = \infty$) and W_T is not bounded either. Thus, the Doob's optional sampling theorem is not valid.